

20. $P(|X| \geq c) = \int_{-\infty}^{-c} f(u)du + \int_c^{+\infty} f(u)du$

$E(X^2) = \int_{-\infty}^{+\infty} u^2 f(u)du$ \rightarrow 被积函数非负

故 $E(X^2) \geq \int_{-c}^{-c} u^2 f(u)du + \int_c^{+\infty} u^2 f(u)du, u^2 \geq c^2$

$\geq \int_{-\infty}^{-c} c^2 f(u)du + \int_c^{+\infty} c^2 f(u)du$
 $= c^2 P(|X| \geq c)$

故 $P(|X| \geq c) \leq \frac{E(X^2)}{c^2}, \forall c > 0$ 成立

21. 记 $Y = |X - m|, Y \geq 0$

① 若 Y 有密度函数 $f_Y(u)$

$P(|X - m| > c) = P(Y > c) = \int_c^{+\infty} f_Y(u)du$

$E(|X - m|) = E(Y) = \int_0^{+\infty} u f_Y(u)du \geq \int_c^{+\infty} u f_Y(u)du \geq \int_c^{+\infty} c f_Y(u)du = c P(Y > c)$

故 $P(Y > c) < \frac{1}{c} E(Y), P(|X - m| > c) < \frac{1}{c} E(|X - m|)$

② 若 Y 是离散的, 没有密度函数

~~$P(Y > c) = \sum_{k > c} P(Y = k)$~~ 记 Y 所有取值构成的集合为 A . A 中大于 c 的所有元素构成 $B, B \subseteq A$

故 $E(Y) = \sum_{k \in A} k P(Y = k) \geq \sum_{k \in B} k P(Y = k) \geq \sum_{k \in B} c P(Y = k) = c \sum_{k \in B} P(Y = k) = c P(Y > c)$

故 $P(Y > c) < \frac{1}{c} E(Y)$

综上, $P(|X - m| > c) < \frac{1}{c} E(|X - m|)$

23. X 分布函数 $F_X(x) = \int_{-\infty}^x \varphi(u)du, x \in \mathbb{R}$

$F_X(x) = P(|X| \leq x) = P(-x \leq X \leq x) = F_X(x) - F_X(-x)$

$= \int_{-\infty}^x \varphi(u)du - \int_{-\infty}^{-x} \varphi(u)du$

故 $|X|$ 的分布函数

$= \int_{-\infty}^x \varphi(u)du - \int (1 - \int_{-x}^{+\infty} \varphi(u)du)$

是 $2\Phi(x) - 1$

$= 2 \int_{-\infty}^x \varphi(u)du - 1$

$= 2 F_X(x) - 1 = 2\Phi - 1, x \geq 0$



25. X 的分布函数 $F_{X^2}(x) = P(X^2 \leq x) = P(-\sqrt{x} \leq X \leq \sqrt{x}) = 2P(X \leq \sqrt{x}) - 1$

$$F_{X^2}(x) = 2 \int_{-\infty}^{\sqrt{x}} \varphi(u) du - 1$$

X 的密度 $f(x) = F'_{X^2}(x) = 2 \cdot \frac{1}{2\sqrt{x}} \cdot \varphi(\sqrt{x}) = \frac{1}{\sqrt{2\pi x}} e^{-\frac{x}{2}}, x \geq 0$

根据密度函数的归一化条件, $\int_0^{+\infty} f(x) dx = \frac{1}{\sqrt{\pi}} \int_0^{+\infty} \frac{1}{\sqrt{2x}} e^{-\frac{x}{2}} dx = 1$

令 $t = \frac{x}{2}$, $dt = \frac{1}{2} dx$

$$\frac{1}{\sqrt{\pi}} \int_0^{+\infty} \frac{1}{\sqrt{2t}} e^{-t} \cdot 2dt = 1 \Rightarrow \int_0^{+\infty} \frac{1}{\sqrt{t}} e^{-t} dt = \sqrt{\pi}$$

29. 记 $Z = |X - Y| \geq 0$

根据 Chebyshev's 不等式, $P(Z > \varepsilon) \leq \frac{1}{\varepsilon^2} E(Z^2) = 0, \forall \varepsilon > 0$ 恒成立

又 $\because P(Z > \varepsilon) \geq 0$

$\therefore P(Z > \varepsilon) = 0, \forall \varepsilon > 0$ 恒成立

根据概率归一化, $P(\Omega) = 1, P(\Omega) = P(Z=0) + P(Z>0) = P(Z=0) = 1$

故 $P(|X - Y| = 0) \Rightarrow P(X = Y) = 1$

30. 令 $X^0 = X - E(X), Y^0 = Y - E(Y), \hat{X} = X^0 / \sigma(X), \hat{Y} = Y^0 / \sigma(Y)$

$$\rho(X, Y) = \frac{E(X^0 Y^0)}{\sqrt{E(X^0)^2} \sqrt{E(Y^0)^2}} = \frac{E(\hat{X} \hat{Y}) \sigma(X) \sigma(Y)}{\sigma(X) \sigma(Y) \sqrt{E(\hat{X}^2)} \sqrt{E(\hat{Y}^2)}} = 1$$

故 $E(\hat{X} \hat{Y}) = \sqrt{E(\hat{X}^2)} \sqrt{E(\hat{Y}^2)}$

$$\sigma^2(\hat{X}) = 1 = E(\hat{X}^2) - E^2(\hat{X}) = E(\hat{X}^2)$$

同理 $E(\hat{Y}^2) = 1$, 故 $E(\hat{X} \hat{Y}) = 1$

$$E(\hat{X} - \hat{Y})^2 = E(\hat{X}^2) + E(\hat{Y}^2) - 2E(\hat{X} \hat{Y}) = 1 + 1 - 2 = 0$$

由 29 可知, $\hat{X} = \hat{Y}$ 恒成立

若 $\rho(X, Y) = -1, E(\hat{X} \hat{Y}) = -\sqrt{E(\hat{X}^2)} \sqrt{E(\hat{Y}^2)} = -1$

$$E(\hat{X} - \hat{Y})^2 = 4 \neq 0$$

结论不成立



4.75

$$(1) \int_{-1.78}^{+0.0} \varphi(u) du = \int_{1.78}^{+0.0} \varphi(u) du = 0.5 - 0.4625 = 0.0375$$

$$(2) \int_{-\infty}^{0.56} \varphi(u) du = \frac{1}{2} \cdot 0.5 + \int_0^{0.56} \varphi(u) du = 0.5 + 0.2123 = 0.7123$$

$$(3) \int_{-1.45}^{+0.0} \varphi(u) du = 0.5 + \int_0^{1.45} \varphi(u) du = 0.5 + 0.4265 = 0.9265$$

$$(4) \int_{2.16}^{+0.0} \varphi(u) du = \frac{0.5}{1} - \int_0^{2.16} \varphi(u) du = 0.5 - 0.4846 = 0.0154$$

$$(5) \int_{-0.80}^{1.53} \varphi(u) du = \int_0^{1.53} \varphi(u) du + \int_0^{0.80} \varphi(u) du = 0.4370 + 0.2881 = 0.7251$$

$$(6) \int_{-\infty}^{-2.52} \varphi(u) du + \int_{1.83}^{+\infty} \varphi(u) du = 1 - \int_0^{2.52} \varphi(u) du - \int_0^{1.83} \varphi(u) du$$

$$= 1 - 0.4941 - 0.4664$$

$$= 0.0395$$

4.78

$$P(Z \geq z_1) = 0.84 \Rightarrow P(Z \leq z_1) = 0.16 \Rightarrow P(0 < Z \leq -z_1) = 0.34$$

$$\text{查表知 } -z_1 \approx 0.995 \quad z_1 = -0.995$$

4.79

$$\mu = 5, \sigma = 2, \text{ 记 } X^* = \frac{X - \mu}{\sigma}, X^* \sim N(0, 1)$$

$$P(X > 8) = P(2X^* + 5 > 8) = P(X^* > 1.5) = 0.5 - 0.4332 = 0.0668$$

4.91

设 X 表示 100 个灯泡中有问题的个数, X 近似服从泊松分布

$$\text{参数 } \lambda = E(X) = 100 \times 3\% = 3$$

$$(1) P(X > 5) = 1 - P(X \leq 5)$$

$$= 1 - \sum_{k=0}^5 \frac{3^k}{k!} e^{-3}$$

$$= 1 - \frac{97}{5} e^{-3}$$

$$(2) P(1 \leq X \leq 3) = P(X=1) + P(X=2) + P(X=3) = 3e^{-3} + \frac{9}{2}e^{-3} + \frac{9}{2}e^{-3} = 12e^{-3}$$

$$(3) P(X \leq 2) = P(X=0) + P(X=1) + P(X=2)$$

$$= e^{-3} + 3e^{-3} + \frac{9}{2}e^{-3}$$

$$= \frac{17}{2}e^{-3}$$

